

life-style risk factors relating to the fact that cannabis consumption is frequently associated with tobacco use and to some festive habits in young males (3). The higher prevalence of intracranial arterial stenosis as a cause of IS in CU in our series is in accordance with the potential involved mechanisms of stroke in CU represented by reversible cerebral vasoconstriction, and impairment of cerebral mitochondrial respiration induced by tetrahydrocannabinol (4,5). The reversibility of vasoconstriction was also described in peripheral arteritis associated with cannabis use (3). Our data demonstrate that a favorable functional capacity is common in young patients suffering from stroke independently of cannabis use that is likely due to age-related enhanced brain plasticity. However, in the whole series, 18% of patients retained significant disability, along with 5 deaths.

Fighting stroke must remain a priority, including in young adults. The first step may be to inform the public regarding the potential occurrence of stroke associated with cannabis use and other life-style risk factors (3), particularly nowadays when cannabis use is encouraged by new legislations worldwide.

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Association of APOC3 Loss-of-Function Mutations With Plasma Lipids and Subclinical Atherosclerosis



The Multi-Ethnic BioImage Study

In 2008, Pollin et al. (1) identified 1 mechanism of lowering triglyceride-rich lipoproteins among the Lancaster Amish, loss of apolipoprotein C-III (APOC3) gene function, which was associated with reduced subclinical atherosclerosis (i.e., coronary arterial calcification on cardiac computed tomography). We recently extended these observations to show that APOC3 loss-of-function mutations also reduce risk for clinical atherosclerotic cardiovascular disease (ASCVD) (2,3). Here, we address 2 questions related to mutations that reduce APOC3 function: 1) do these mutations also associate with lower plasma low-density lipoprotein cholesterol (LDL-C); and 2) do these mutations reduce subclinical atherosclerosis in the general population, particularly in individuals with ancestry outside of Europe?

We studied 6,699 individuals of European, African, Asian, and Hispanic ancestries from the BioImage Study (BioImage Study: A Clinical Study of Burden of Atherosclerotic Disease in an At-Risk Population, [NCT00738725](https://clinicaltrials.gov/ct2/show/study/NCT00738725)), a prospective, observational study aimed at characterizing subclinical atherosclerosis in U.S. adults (55 to 80 years old) at risk for clinical ASCVD (4). We genotyped 4 APOC3 (NM_000040.1) loss-of-function mutations (IVS2+1G→A, A43T, R19X, IVS3+1G→T) using the Illumina HumanExome Bead-Chip Array v1.1 (Illumina, San Diego, California). Written informed consent was obtained from all study participants according to a protocol approved by the Western Institutional Review Board, Olympia, Washington. Fasting blood lipids were measured at the baseline examination. Blood lipids were adjusted for the presence of statin medications to reflect the observation that statins, on average, reduce total cholesterol by 20% and LDL-C by 30% (5). Noninvasive assessments for subclinical atherosclerosis (coronary arterial calcification [CAC], carotid plaque, and carotid intima media thickness [CIMT]) were performed at the baseline examination on a mobile imaging facility as previously described (4).

A total of 6,395 subjects passed all quality-control measures. Variant calling was performed using

TABLE 1 Association of *APOC3* Loss-of-Function Mutation Carrier Status With Blood Lipid Levels and Subclinical Atherosclerosis

	Noncarriers (n = 6,331)	Carriers (n = 64)	Effect Estimate	95% CI	p Value
Blood lipids					
Triglycerides, mg/dl*	166.1 ± 96.2	91.2 ± 44.1	−43.7%	−36.6 to −49.9	1.83×10^{-21}
HDL cholesterol, mg/dl	55.6 ± 15.2	66.6 ± 15.0	+11.1	7.6 to 14.6	3.55×10^{-10}
LDL cholesterol, mg/dl	130.3 ± 36.4	131.6 ± 35.8	+1.5	−7.4 to 10.3	0.75
Subclinical atherosclerosis					
Coronary arterial calcification, Agatston units†	46.0 (0.0 to 245.0)	29.0 (0.0 to 227.5)	−27.9	−51.1 to −4.7	0.019
Carotid plaque, mm ² †	183.8 (0.0 to 555.9)	112.8 (0.0 to 367.2)	−8.7	−73.5 to 56.0	0.79
Carotid intima media thickness, mm*	0.76 ± 0.16	0.74 ± 0.13	−1.7%	−6.3 to 3.0	0.47

Values are represented as effect estimates (95% confidence intervals) between *APOC3* loss-of-function mutation carriers and noncarriers after adjustment for covariates. Covariates used were age, sex, ancestry, and principal components of ancestry. *Effect estimates for triglycerides and carotid intima media thickness were derived from natural log transformation and are expressed as percent difference in carriers compared with noncarriers. †Effect estimates for coronary arterial calcification and carotid plaque derived from median quantile regression and represent differences in medians. Therefore, summary statistics of the distributions in noncarriers and carriers are represented as median (interquartile range).

CI = confidence interval; HDL = high-density lipoprotein; LDL = low-density lipoprotein.

GenCall (Illumina) and zCall (5). Sixty-four heterozygous carriers of *APOC3* loss-of-function mutations (25 IVS2+1G→A, 25 A43T, 13 R19X, 1 IVS3+1G→T; combined minor allele frequency of 0.5%) were identified. Principal components were derived from a set of high-quality, independent variants on the genotyping array using Eigenstrat as has previously been done (3,5). To minimize confounding from systematic differences in allele frequencies by trait, we reduced the observed genetic variation to the top eigenvectors derived from the sample covariance matrix. To test the association of *APOC3* loss-of-function mutation with an outcome, linear regression was used for triglycerides, LDL-C, and high-density lipoprotein cholesterol (HDL-C), and CIMT; triglycerides and CIMT were natural log-transformed. And given the bimodal, skewed distributions of CAC (primary outcome) and carotid plaque, median quantile regression was used for these 2 variables. Age, sex, ethnicity, and principal components of ancestry were used as covariates in all analyses. Given a 2-sided alpha threshold of 0.05, we have >80% power to detect an effect size of 0.16% of variance explained for analyzed traits.

Among carriers and noncarriers of the *APOC3* loss-of-function mutations, there were no significant differences in age, sex, hypertension, diabetes mellitus, body mass index, current smoking, aspirin use, or statin use. There were no significant differences in proportions of carriers amongst each ethnicity group ($p > 0.20$). We replicated the finding that *APOC3* loss-of-function mutations were associated with reduced triglycerides (−43.7%; $p = 1.83 \times 10^{-21}$) and increased HDL-C (11.1 mg/dl;

$p = 3.55 \times 10^{-10}$), with a larger standardized effect on triglycerides compared with HDL-C (−1.17 SD vs. +0.73 SD). When accounting for statin treatment, carriers did not have different LDL-C concentrations compared with noncarriers ($p = 0.75$).

APOC3 loss-of-function mutations were associated with decreased median CAC score (−27.9 U; 95% confidence interval [CI]: −51.08 to −4.67; $p = 0.019$) across all phenotyped participants ($n = 5,631$); this effect was consistent in those of European ancestry (−27.5 U; 95% CI: −67.1 to 12.1) and of non-European ancestry (−5.62 U; 95% CI: −39.2 to 27.9). Neither carotid plaque ($p = 0.79$) nor CIMT ($p = 0.47$) ($n = 5,746$) differed between carriers and noncarriers (Table 1).

In a multiethnic study of U.S. adults, *APOC3* loss-of-function mutation carriers had reduced plasma triglycerides, higher HDL-C, and a decreased burden of coronary arterial calcification. These data support the notion that *APOC3* deficiency reduces coronary atherosclerosis in the general population. Whether pharmacological inhibition of *APOC3* will reduce ASCVD risk remains to be tested.

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Analysis of Dual Antiplatelet Therapy



In a post-hoc analysis, Yeh et al. (1) provide evidence that dual-antiplatelet therapy with thienopyridine and aspirin, for 30 (vs. 12) months after coronary

stenting, had beneficial effects regardless of whether patients had acute myocardial infarction (MI) or less acute presentations. In particular, stent thrombosis was reduced—although bleeding was increased—similarly in both groups. Clinicians should, however, consider various elements of the study design when interpreting results and using the findings in clinical practice.

Importantly, the study classified patients with unstable angina as having less acute (“no MI”) presentations. This subset accounted for nearly one-fourth (22.6%) of the 8,072 patients without MI, or roughly 1,800 patients—representing more than one-half of the 3,576 patients with MI. Combining unstable angina patients within the non-MI group is of more utility in allowing for conclusions to be drawn on the MI group, given that the non-MI group represents a population of patients with a wide range of risk of subsequent events—and it is unclear whether the beneficial effects of prolonged treatment was driven by those at higher risk within this group. Perhaps a sensitivity analysis could be done, excluding patients with unstable angina from the non-MI group, to see if the overall results are affected.

From a more general perspective, the reader may not immediately recognize that the current analyses represent a post-hoc subgroup analysis of a previously completed randomized trial, with randomization 12 months after initial enrollment. We would defend the authors from reflexive criticisms that such analyses are always problematic, but we also note that whether the findings are valid, or due to chance or bias, is always a consideration. For example, lacking a formal power calculation linked to antecedent MI status, the discordant findings of similarly reduced stent thrombosis without an interaction, compared to differential reduction of major adverse cardiovascular and cerebrovascular events with an interaction, have uncertain implications.

Finally, as a specific methodological issue, the data in Figure 1 in the paper by Yeh et al. (1) indicate that (25,682 total minus 5,844 not eligible equals) 19,838 patients were eligible for randomization, but 8,190 were not randomized at 12 months for the reasons stated. Given that this group represents (8,190 divided by 19,838 equals) 41.3% of eligible patients, it would be useful to know whether pertinent baseline characteristics (including the distribution of patients in MI and non-MI groups) differ between analyzed and excluded patients, for better understanding the overall results.